



Subset Gaussian
Cloning and
Connections With
High Dimensional
RRT.

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June 15, 2025



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n participants, m questions.

Warner's additive model outputs the matrix,

	P_1	P_2	$\dots$	P_{m-1}	P_m
Q_1	0.12	-1.05	$\dots$	-0.43	0.27
Q_2	-0.78	0.46	$\dots$	0.55	-0.31
$\vdots$	$\vdots$	$\vdots$	$\ddots$	$\vdots$	$\vdots$
Q_{m-1}	-0.24	0.89	$\dots$	0.07	-0.19
Q_m	0.45	-0.67	$\dots$	0.78	0.11

Question vector Q and response vector P .

Adding structure to the signal!

Let $P = \mu \in \{-1, 0, 1\}^n$ with $\|\mu\|_0 = k$.

- ▶ ± 1 are two differing perspectives.
- ▶ 0 is a “moderate” or “ambivalent” perspective with respect to ± 1

Which $Q \in \mathbb{R}^n$ vector to consider?
 $Q = \mathbf{1}$ is equivalent to:

“Do you agree with perspective 1?”

“1’s” $\rightarrow 1$, “-1’s” $\rightarrow -1$ and “0’s” $\rightarrow 0$

Question vector = $\mathbf{1}$,

Response vector = μ ,

Warner's model becomes,

$$Y = \lambda \frac{\mathbf{1}\mu^\top}{n^{1/2}k^{1/2}} + G$$

Question vector = $\mathbf{1}$,

Response vector = μ ,

Warner's model becomes,

$$Y = \underbrace{\lambda}_{\text{SNR}} \underbrace{\frac{\mathbf{1}\mu^\top}{k^{1/2}n^{1/2}}}_{\text{Rescaled Signal}} + \underbrace{G}_{\text{Warner's Additive Noise}},$$

$$\|\mu\|_0 = k \text{ and } G_{i,j} \sim \mathcal{N}(0, 1)$$

Instead of $\mathbf{1}$, consider $\mathbf{q} \in \{-1, 0, 1\}^m$
(possible random) with $\|\mathbf{q}\|_0 = \ell$.

Warner's matrix $Y = Y_{i,j} = q_i \mu_j$.

Re-normalizing,

$$Y = \lambda \frac{\mathbf{q} \mu^\top}{\ell^{1/2} k^{1/2}} + G$$

Let $i \in [m]$ and $j \in [n]$.

$$q_i \neq \mu_j \neq 0 \text{ gives } Y_{i,j} = \frac{\lambda}{\sqrt{\ell_k}} + G_{i,j}.$$

$$q_i = \mu_j \neq 0 \text{ gives } Y_{i,j} = \frac{-\lambda}{\sqrt{\ell_k}} + G_{i,j}.$$

$$q_i = 0 \text{ or } \mu_j = 0 \text{ gives } Y_{i,j} = G_{i,j}.$$

$$\begin{array}{ccc}
Y & = & \lambda \frac{q\mu^\top}{\ell^{1/2}k^{1/2}} + G \\
(m = n) & & \downarrow \quad (?) \\
Y & = & \lambda \frac{\mu\mu^\top}{k^{1/2}k^{1/2}} + G \\
& & \downarrow \quad (tensor) \\
Y & = & \lambda \frac{\mu^{\otimes r}}{k^{r/2}} + G
\end{array}$$

In what follows, λ large allows exact recover of μ from Warner's matrix.

Why is this problematic in RRT?

We turn to the model,

$$Y = \lambda \frac{\mu^{\otimes r}}{k^{r/2}} + G,$$

$$Y \in \mathbb{R}^{n^{\otimes r}}, \lambda_n \geq 0, G_{i_1, \dots, i_r} \sim \mathcal{N}(0, 1), \\ \mu \in \{-1, 0, 1\}^n \text{ and } \|\mu\|_0 = k$$

- ▶ *Many applications:* tensor PCA, compressed sensing, community detection.
- ▶ **More Importantly**, in this model exhibits a stat-comp-local gap.

First, we look at STAT.

Question?

Let ν be uniform over
 $\{v \in \{-1, 0, 1\}^n : \|v\|_0 = k\}$ be the prior.

What is the optimal (in terms of exact recovery) statistical estimator of μ .



Maximum a posteriori, i.e.

$$\hat{\mu} = \arg \max_{\mu' \in \mu} \langle Y, (\mu')^{\otimes r} \rangle.$$

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For our problem of focus,

$$\text{opt}_{\text{STAT}}(\lambda) = \Theta(\sqrt{k} p \log(n))$$

and

$$\text{opt}_{\text{COMP}}(\lambda) = \Theta((k^{r/2} \wedge n^{r/4}) p \log(n)).$$

A Comp-Local gap occurs when the best known (non-“local”) poly-time algorithm requires SNR λ' , but there is evidence that a “local” poly-time algorithm requires $\tilde{\lambda} = \omega(\lambda')$.

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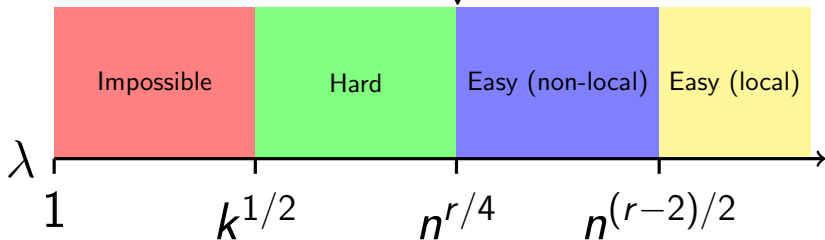
Recalling the previous slide,

$$\text{opt}_{\text{COMP}}(\lambda) = (k^{r/2} \wedge n^{r/4})p\text{-log}(n)$$

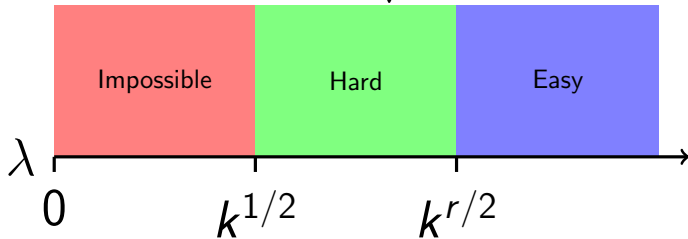
and

$$\text{opt}_{\text{LOCAL}} = (n^{(r-2)/2})p\text{-log}(n).$$

If $k \geq \sqrt{n}$,



If $k < \sqrt{n}$,



Generating correlated noise,

Require: $M \in \mathbb{N}$

- 1: **for** $i = 1$ to n **do**
- 2: Sample $G_i^1, \dots, G_i^M$ i.i.d. from $\mathcal{N}(0, M)$.
- 3: $\bar{G}_i \leftarrow \frac{1}{M} \sum_{j=1}^M G_i^j$.
- 4: **for** $t = 1$ to M **do**
- 5: $Z_{i,t} \leftarrow G_i^t - \bar{G}_i$.
- 6: **end for**
- 7: **end for**
- 8: **return** $((Z_i^t)_{i \in [n], t \in [M]})$

Randomized Greedy with random thresholds,

Require: $Y \in \mathbb{R}^{n \otimes r}$, $S_0 \in \{-1, 0, 1\}^n$, $\gamma \geq 0$, $M \in \mathbb{N}$

- 1: $t \leftarrow 0$ and $t_i \leftarrow 0$ for each $i \in [n]$.
- 2: Let $((Z_i^t)_{i \in [n], t \in [M]})$ be the output of the above algorithm with input M
- 3: **while** $t \leq M$ **do**
- 4: Let $\mathcal{N}(S_t) = \{\sigma' \in \{-1, 0, 1\}^n \mid d_H(S_t, \sigma') = 1\}$
- 5: Choose a random element $\sigma' \in \mathcal{N}(S_t)$ uniformly
- 6: Let $i \in [n]$ be such that $\sigma'_i \neq (S_t)_i$.
- 7: $t_i \leftarrow t_i + 1$, $t \leftarrow t + 1$
- 8: **if** $H_{\frac{r+1}{2}, \gamma}(\sigma') - H_{\frac{r+1}{2}, \gamma}(S_t) > 0$
- 9: $S_{t+1} \leftarrow \sigma'$
- 10: **end if**
- 11: **end while**
- 12: **return** S_{M+1}

Above $H_{\frac{r+1}{2}, \gamma}(S) = \langle S, Y \rangle - \gamma \|S\|_0^{\frac{r+1}{2}}$.

For simplicity, $\lambda = \Omega(\sqrt{k}n^{(r-1)/4})$.

(1) Justify the following heuristic: If $\cos(\mu, S_1)^{r-1} \geq \frac{\sqrt{k}}{\lambda} \text{p-log}(n)$, then our algorithm recovers μ .

(2) Find an initialization where $\cos(\mu, S_1) = n^{-1/4}$.

(3) By rearranging the inequality in (1), we then have $\lambda \geq \sqrt{k} n^{\frac{r-1}{4}} \text{p-log}(n)$, as desired.

Parameterize the Hamming distance one transition by choosing a coordinate p_t of S_t to change to $q_t \neq S_t$.

Marginally,

$$\begin{aligned}
 H_{\frac{r+1}{2}, \gamma}(S') - H_{\frac{r+1}{2}, \gamma}(S_t) &\approx \\
 (q_t - (S_t)_{p_t}) \mu_{p_t} \frac{\lambda}{k^{r/2}} \langle \theta, S_t \rangle^{r-1} &+ \langle S_t, G \rangle - (|q_t| - |S_{p_t}^t|) \gamma \|S^t\|_0^{(r-1)/2} \\
 = \frac{\lambda}{\sqrt{k}} (q_t - (S_t)_{p_t}) \mu_{p_t} \cos(S_t, \theta)^{r-1} &- (|q_t| - |S_{p_t}^t|) \gamma + Z_t
 \end{aligned}$$

So why not bound Z_t and do case work?

Why can't you simply bound Z_t ?

- ▶ First, the random variables Z_t and $Z_{t'}$ are correlated as S_t is close to $S_{t'}$.
- ▶ Second, we have to bound Z_t over all possible paths of this walk. If we have made n transitions, this set of possible paths grow exponentially.

Alternately, we could use sub-set Gaussian cloning to establish $\lambda \geq \sqrt{k} n^{\frac{r-1}{4}} \text{p-log}(n)$.

Intuition

- ▶ Many algorithms reuse noise matrix G , causing correlations between steps.
- ▶ To avoid the complexity, we could assume: $\tilde{Y}_t = \frac{\lambda}{k^{r/2}} \theta^{\otimes r} + \tilde{G}_t$ with $\tilde{G}_t \sim \mathcal{N}(0, \cdot)$, for each step t .
- ▶ Solution: design an algorithm that simulates fresh-noise behavior.

Spot the step which simulates fresh noise?

Require: $Y \in \mathbb{R}^{n^{\otimes r}}$, $S_0 \in \{-1, 0, 1\}^n$, $\gamma \geq 0$, $M \in \mathbb{N}$

- 1: $t \leftarrow 0$ and $t_i \leftarrow 0$ for each $i \in [n]$.
- 2: Let $((Z_i^t)_{i \in [n], t \in [M]})$ be the output of the above algorithm with input M
- 3: **while** $t \leq M$ **do**
- 4: Let $\mathcal{N}(S_t) = \{\sigma' \in \{-1, 0, 1\}^n \mid d_H(S_t, \sigma') = 1\}$
- 5: Choose a random element $\sigma' \in \mathcal{N}(S_t)$ uniformly
- 6: Let $i \in [n]$ be such that $\sigma'_i \neq (S_t)_i$.
- 7: $t_i \leftarrow t_i + 1$, $t \leftarrow t + 1$
- 8: **if** $H_{\frac{r+1}{2}, \gamma}(\sigma') - H_{\frac{r+1}{2}, \gamma}(S_t) > \|(\sigma')^{\otimes r} - (S_t)^{\otimes r}\|_F Z_i^{t_i}$ **then**
- 9: $S_{t+1} \leftarrow \sigma'$
- 10: **end if**
- 11: **end while**
- 12: **return** S_{M+1}

General setup

- ▶ Model: $Y_i = \mu_i^* + Z_i$, with $Z_i \sim \mathcal{N}(0, 1)$.
- ▶ Input: subsets $\Delta_t \subseteq [M]$, and a total number of draws M .
- ▶ Goal: output $X_{\Delta_t} \sim \mu_{\Delta_t}^* + \mathcal{N}(0, M \cdot I)$, independently across t .

Input: $Y \in \mathbb{R}^N$, subsets Δ_t , usage cap M .

1. For each $i \in [N]$, sample $G_i^1, \dots, G_i^M \sim \mathcal{N}(0, M)$.
2. Compute $\bar{G}_i = \frac{1}{M} \sum_{j=1}^M G_i^j$.
3. Initialize $b_i^0 = 0$, track how many times i is used.

4. At step t :

For each $i \in \Delta_t$, set $b_i^t = b_i^{t-1} + 1$

Otherwise, set $b_i^t = b_i^{t-1}$

$$X_{\Delta_t} = Y_{\Delta_t} + \left(G_{\Delta_t}^{b^t} - \bar{G}_{\Delta_t} \right)$$

5. Stop when any element is used $> M$ times (i.e. we can't inject any more fresh noise). Denote this time as T_M .

Theorem: For all $t \in [T_M]$,

$$X_{\Delta_t} \stackrel{d}{=} \mu_{\Delta_t}^* + Z'_{\Delta_t}, \quad Z'_{\Delta_t} \sim \mathcal{N}(0, M \text{Id}_{|\Delta_t|})$$

- ▶ Each X_{Δ_t} is independent of the others.
- ▶ Outputs simulate fresh Gaussian noise across steps.
- ▶ Cost: variance of noise is inflated by M .

$$\lambda\sqrt{k}(q_t-(S_t)p_t)\mu_{p_t}\cos(S_t,\theta)^{r-1}-(|q_t|-|S_{p_t}^t|)\gamma+Z_t'\geq 0$$

Standard asymptotic theory gives

$$\max_t |Z_t'| \leq 2\sqrt{M\log(n)} \text{ wp } 1 - o(1),$$

$$\lambda\sqrt{k}(q_t-(S_t)p_t)\mu_{p_t}\cos(S_t,\theta)^{r-1}\geq \pm 2\sqrt{M\log(n)}+\gamma(|q_t|-|S_{p_t}^t|).$$

$$\equiv (q_t-(S_t)p_t)\mu_{p_t}\cos(S_t,\theta)^{r-1}\geq \frac{\sqrt{k}}{\lambda}\left(\pm 2\sqrt{M\log(n)}+\gamma(|q_t|-|S_{p_t}^t|)\right).$$

We then can rigorously prove condition:

$$\cos(S_t,\theta)^{r-1}\geq \frac{\sqrt{k}}{\lambda}\text{p-log}(n).$$

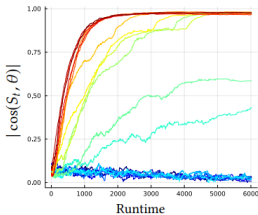


Figure 3: Randomized greedy for Sparse 3-Tensor PCA when $\theta \in \{-1, 0, 1\}^n$. Mean absolute angle vs time for $\lambda = n^\alpha$, initialized at a uniform random trinary vector.

We predict that $\alpha = 1.4$ is the threshold for fast recovery. Here $n = 150$, $k = 56 \approx n^{0.8}$, and $\gamma = \log n$.

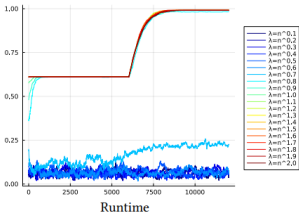


Figure 5: Two-stage algorithm for sparse 3-tensor PCA when $\theta \in \{-1, 0, 1\}^n$. Mean absolute angle vs. time for $\lambda = n^\alpha$, initialised at S_{HOM} . We predict that $\alpha = 0.75$ is the threshold for fast recovery. Here $n = 150$, $k = 56 \approx n^{0.8}$, and $\gamma = \log n$.

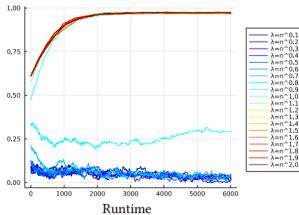


Figure 4: Randomized greedy for Sparse 3-Tensor PCA when $\theta \in \{-1, 0, 1\}^n$. Mean absolute angle vs time for $\lambda = n^\alpha$, initialized at S_{HOM} . We predict that $\alpha = 0.9$ is the threshold for fast recovery. Here $n = 150$, $k = 56 \approx n^{0.8}$, and $\gamma = \log n$.

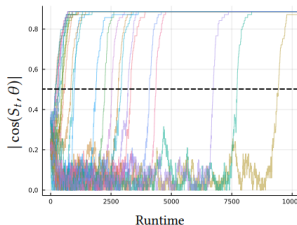


Figure 6: Stage one of the two-stage algorithm for sparse 3-tensor PCA when $\theta \in \{-1, 0, 1\}^n$. Absolute angle vs. time for $\lambda = \Theta(n^{3/4})$, initialised at S_{HOM} .

Here $n = 150$, $k = 116 \approx n^{0.95}$, and we plot 29 simulations (filtered from 400 total simulations) which exhibit the most visible oscillatory phase. The dashed black line (where $|\cos(S_t, \theta)| \approx .502$) is a prediction for when this stage transitions from an oscillatory phase (below the line) to a monotonically increasing phase (above the line). Color is only used for visual clarity.

- ▶ Rigorous analysis of iterative algorithms.
- ▶ Variance inflation is mild under $M = p \log(n)$.
- ▶ Takeaway: **Subset Gaussian Cloning bridges algorithm design and clean probabilistic models.**

- ▶ Non-Gaussian
- ▶ More general algorithms
- ▶ A more convenient noising scheme

